

# Quantum Algorithms for the Pathwise Lasso

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## High dimensional regression

$$Y = X\beta + \epsilon$$

here  $X$  is  $n \times d$  matrix, where  $n \ll d$ . More features, assets than observations. Thus traditional OLS solution  $\beta = (X^T X)^{-1} X^T Y$  is not feasible. Sparsity is thus desired:  $(X^T X)$  is rank deficient.

# What is Lasso?



$\beta$



# Regressor Selection

$$\tilde{\beta} = \underset{\beta \in \mathbb{R}^d \setminus \{0\}}{\operatorname{argmin}} \|Y - X\beta\|_2^2 + \lambda \|\beta\|_1 \quad \text{for all } \lambda > 0.$$

We will use this set up. There are two cases to consider.

1. We seek a sparse portfolio to emulate  $Y$  such as in index tracking. The  $\ell_1$  penalty captures the transaction costs.
2. In an mean-variance portfolio, we can set  $Y = 0$ , thus our problem becomes

$$\tilde{\beta} = \underset{\beta \in \mathbb{R}^d \setminus \{0\}}{\operatorname{argmin}} \|X\beta\|_2^2 + \lambda \|\beta\|_1 \quad \text{for all } \lambda > 0.$$

This is the same as a standard mean-variance problem without the constraint of summing to 1.

# Portfolio Selection - Sparse

We desire a sparse solution.

1. Not all assets are included in the portfolio so  $\beta_{\mathcal{I}} = 0$  for some inactive set of assets.
2. Introducing the  $\ell_1$  norm induces sparsity but we are still left with the problem of choosing  $\lambda$
3. Perhaps the answer is homotopy methods, where we seek a regularized path  $\beta(\lambda)$ . So the weights are a function of varying  $\lambda$  allowing for greater flexibility in choice.
4. LARS or Homotopy algorithm method gives us such a regularized path.

# Classical Algorithm

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**Algorithm 1:** Classical LARS algorithm for the pathwise Lasso

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**Input:** Vector  $\mathbf{y} \in \mathbb{R}^n$  and matrix  $\mathbf{X} \in \mathbb{R}^{n \times d}$

**Output:** Exact regularisation path  $\mathcal{P}$

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1 Initialise  $\mathcal{A} = \operatorname{argmax}_{j \in [d]} |\mathbf{X}_j^\top \mathbf{y}|$ ,  $\mathcal{I} = [d] \setminus \mathcal{A}$ ,  $\lambda_0 = \|\mathbf{X}^\top \mathbf{y}\|_\infty$ ,  $\hat{\boldsymbol{\beta}}(\lambda_0) = \mathbf{0}$ ,  $t = 0$ 
2 while  $\mathcal{I} \neq \emptyset$  do
3    $\boldsymbol{\eta}_{\mathcal{A}} \leftarrow \operatorname{sign}(\mathbf{X}_{\mathcal{A}}^\top (\mathbf{y} - \mathbf{X}_{\mathcal{A}} \hat{\boldsymbol{\beta}}_{\mathcal{A}}(\lambda_t)))$ 
4    $\boldsymbol{\mu} \leftarrow \mathbf{X}_{\mathcal{A}}^+ \mathbf{y}$  //  $\boldsymbol{\mu} \in \mathbb{R}^{|\mathcal{A}|}$ 
5    $\boldsymbol{\theta} \leftarrow (\mathbf{X}_{\mathcal{A}}^\top \mathbf{X}_{\mathcal{A}})^+ \boldsymbol{\eta}_{\mathcal{A}}$  //  $\boldsymbol{\theta} \in \mathbb{R}^{|\mathcal{A}|}$ 
6    $\Lambda_i^{\text{cross}} \leftarrow \frac{\mu_i}{\theta_i} \cdot \mathbf{1} \left[ \frac{\mu_i}{\theta_i} \leq \lambda_t \right] \quad \forall i \in \mathcal{A}$ 
7    $\Lambda_i^{\text{join}} \leftarrow \frac{\mathbf{X}_i^\top (\mathbf{y} - \mathbf{X}_{\mathcal{A}} \boldsymbol{\mu})}{\pm 1 - \mathbf{X}_i^\top \mathbf{X}_{\mathcal{A}} \boldsymbol{\theta}} \quad \forall i \in \mathcal{I}$ 
8    $i_{t+1}^{\text{cross}} \leftarrow \operatorname{argmax}_{i \in \mathcal{A}} \{\Lambda_i^{\text{cross}}\}$  and  $\lambda_{t+1}^{\text{cross}} \leftarrow \max_{i \in \mathcal{A}} \{\Lambda_i^{\text{cross}}\}$ 
9    $i_{t+1}^{\text{join}} \leftarrow \operatorname{argmax}_{i \in \mathcal{I}} \{\Lambda_i^{\text{join}}\}$  and  $\lambda_{t+1}^{\text{join}} \leftarrow \max_{i \in \mathcal{I}} \{\Lambda_i^{\text{join}}\}$ 
10   $\lambda_{t+1} \leftarrow \max\{\lambda_{t+1}^{\text{join}}, \lambda_{t+1}^{\text{cross}}\}$ 
11   $\hat{\boldsymbol{\beta}}_{\mathcal{A}}(\lambda_{t+1}) \leftarrow \boldsymbol{\mu} - \lambda_{t+1} \boldsymbol{\theta}$ 
12  if  $\lambda_{t+1} = \lambda_{t+1}^{\text{cross}}$  then
13    | Move  $i_{t+1}^{\text{cross}}$  from  $\mathcal{A}$  to  $\mathcal{I}$ 
14  else
15    | Move  $i_{t+1}^{\text{join}}$  from  $\mathcal{I}$  to  $\mathcal{A}$ 
16   $t \leftarrow t + 1$ 
17 return coefficients  $[(\lambda_0, \hat{\boldsymbol{\beta}}(\lambda_0)), (\lambda_1, \hat{\boldsymbol{\beta}}(\lambda_1)), \dots]$ 
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# Why Quantum

Evaluating a regularized path is highly non-trivial and in fact an NP-hard problem. How to determine  $\tau$  for the the next  $\lambda$ .  
Quantizing parts of the classical algorithm may help speed up.  
As complexity of the classical algorithm is  
We will analyse this algorithm and its variants.



## How our work is new and different

Evaluating a regularized path is highly non-trivial and in fact an NP-hard problem. How to determine  $\tau$  for the the next  $\lambda$ .

Existing work focuses on a fixed  $\lambda$

$$\tilde{\beta} = \operatorname{argmin}_{\beta \in \mathbb{R}^d} \|Y - X\beta\|_2^2 + \lambda \|\beta\|_1 \quad \text{fixed } \lambda > 0.$$

### Definition

Given  $\epsilon \geq 0$ , we say that a vector  $\tilde{\beta} \in \mathbb{R}^d$  is an approximate Lasso solution with error  $\epsilon(\lambda)$  with  $\lambda > 0$  if

$$\frac{1}{2} \|\mathbf{y} - \mathbf{X}\tilde{\beta}\|_2^2 + \lambda \|\tilde{\beta}\|_1 - \left( \min_{\beta \in \mathbb{R}^d} \frac{1}{2} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \lambda \|\beta\|_1 \right) \leq \epsilon(\lambda)$$

# Approximate quantum LARS

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**Algorithm 3:** Approximate quantum LARS algorithm for the pathwise Lasso

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**Input:**  $T \in \mathbb{N}$ ,  $\delta \in (0, 1)$ ,  $\epsilon > 0$ ,  $\mathbf{y} \in \mathbb{R}^n$ ,  $\mathbf{X} \in \mathbb{R}^{n \times d}$ , and  $\{\alpha_{\mathcal{A}} \geq \|\mathbf{X}_{\mathcal{A}}^+ \mathbf{X}_{\mathcal{A}^c}\|_1\}_{\mathcal{A} \subseteq [d]: |\mathcal{A}| \leq T}$

**Output:** Regularisation path  $\tilde{\beta}$  with  $T$  kinks and error  $\lambda \epsilon \|\tilde{\beta}(\lambda)\|_1$  with probability  $\geq 1 - \delta$

1 Initialise  $\mathcal{A} = \operatorname{argmax}_{j \in [d]} |\mathbf{X}_j^\top \mathbf{y}|$ ,  $\mathcal{I} = [d] \setminus \mathcal{A}$ ,  $\lambda_0 = \|\mathbf{X}^\top \mathbf{y}\|_\infty$ ,  $\tilde{\beta}(\lambda_0) = 0$ ,  $t = 0$

2 **while**  $\mathcal{I} \neq \emptyset$ ,  $t \leq T$  **do**

3      $\boldsymbol{\eta}_{\mathcal{A}} \leftarrow \frac{1}{\lambda_t} \mathbf{X}_{\mathcal{A}}^\top (\mathbf{y} - \mathbf{X}_{\mathcal{A}} \tilde{\beta}_{\mathcal{A}}(\lambda_t))$

4      $\boldsymbol{\mu} \leftarrow \mathbf{X}_{\mathcal{A}}^+ \mathbf{y}$  //  $\boldsymbol{\mu} \in \mathbb{R}^{|\mathcal{A}|}$

5      $\boldsymbol{\theta} \leftarrow (\mathbf{X}_{\mathcal{A}}^\top \mathbf{X}_{\mathcal{A}})^+ \boldsymbol{\eta}_{\mathcal{A}}$  //  $\boldsymbol{\theta} \in \mathbb{R}^{|\mathcal{A}|}$

6     Classically compute  $\mathbf{y} - \mathbf{X}_{\mathcal{A}} \boldsymbol{\mu}$  and  $\mathbf{X}_{\mathcal{A}} \boldsymbol{\theta}$  and input them into QRAMs and KP-trees

7     Define  $\Lambda_i^{\text{join}} \triangleq \frac{\mathbf{X}_i^\top (\mathbf{y} - \mathbf{X}_{\mathcal{A}} \boldsymbol{\mu})}{\pm 1 - \mathbf{X}_i^\top \mathbf{X}_{\mathcal{A}} \boldsymbol{\theta}}$  and  $\Lambda_i^{\text{cross}} \triangleq \frac{\mu_i}{\theta_i} \cdot \mathbf{1} \left[ \frac{\mu_i}{\theta_i} \leq \lambda_t \right]$

8      $i_{t+1}^{\text{cross}} \leftarrow \operatorname{argmax}_{i \in \mathcal{A}} \{\Lambda_i^{\text{cross}}\}$  and  $\lambda_{t+1}^{\text{cross}} \leftarrow \max_{i \in \mathcal{A}} \{\Lambda_i^{\text{cross}}\}$

9     Obtain  $\tilde{i}_{t+1}^{\text{join}} \in \{j \in \mathcal{I} : \Lambda_j^{\text{join}} \geq (1 - \frac{\epsilon}{1 + \alpha_{\mathcal{A}}}) \max_{i \in \mathcal{I}} \{\Lambda_i^{\text{join}}\}\}$  with failure probability  $\frac{\delta}{T}$

    (Fact 14 and Lemma 16)

10      $\tilde{\lambda}_{t+1}^{\text{join}} \leftarrow (1 - \frac{\epsilon}{1 + \alpha_{\mathcal{A}}})^{-1} \Lambda_{\tilde{i}_{t+1}^{\text{join}}}^{\text{join}}$

11      $\lambda_{t+1} \leftarrow \min\{\lambda_t, \max\{\lambda_{t+1}^{\text{cross}}, \tilde{\lambda}_{t+1}^{\text{join}}\}\}$

12      $\tilde{\beta}_{\mathcal{A}}(\lambda_{t+1}) \leftarrow \boldsymbol{\mu} - \lambda_{t+1} \boldsymbol{\theta}$

13     **if**  $\lambda_{t+1} = \lambda_{t+1}^{\text{cross}}$  **then**

14         | Move  $i_{t+1}^{\text{cross}}$  from  $\mathcal{A}$  to  $\mathcal{I}$

15     **else**

16         | Move  $\tilde{i}_{t+1}^{\text{join}}$  from  $\mathcal{I}$  to  $\mathcal{A}$

17      $t \leftarrow t + 1$

18 **return** coefficients  $[(\lambda_0, \tilde{\beta}(\lambda_0)), (\lambda_1, \tilde{\beta}(\lambda_1)), \dots]$

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# Lasso Complexity

**Lemma 3** ([10]). *For any  $Y$  and  $X$ , the Lasso path algorithm performs at most*

$$\sum_{k=0}^d \binom{d}{k} 2^k = 3^d$$

*iterations before termination.*

*Proof.* Actually, we give a simple proof. This is just a simple application of the Binomial theorem

$$\sum_{k=0}^d \binom{d}{k} x^k = (1+x)^d$$

where  $x = 2$ . Hence, the identity holds. □

This can be reduced by putting a hard constraint on the number of regressors  $m < \min(n, d)$  but we still have an exponential worst-case complexity.

# Some Bounds - Fixed Design Matrix

Algorithm	Error	Time complexity per iteration
Classical LARS 1	None	$O(n \mathcal{I}  + n \mathcal{A}  +  \mathcal{A} ^2)$
Simple quantum LARS 2	None	$\tilde{O}(n\sqrt{ \mathcal{I} } + n \mathcal{A}  +  \mathcal{A} ^2)$
Approximate classical LARS 4	$\lambda\epsilon\ \tilde{\beta}\ _1$	$\tilde{O}\left(\frac{\gamma_{\mathcal{A}}^{-2} + n\ \mathbf{X}\ _{\max}^2\ \mathbf{X}_{\mathcal{A}}^{\dagger}\ _2^2}{(1-\alpha_{\mathcal{A}})^2\epsilon^2} \mathcal{I}  + n \mathcal{A}  +  \mathcal{A} ^2\right)$
Approximate quantum LARS 3	$\lambda\epsilon\ \tilde{\beta}\ _1$	$\tilde{O}\left(\frac{\gamma_{\mathcal{A}}^{-1} + \sqrt{n}\ \mathbf{X}\ _{\max}\ \mathbf{X}_{\mathcal{A}}^{\dagger}\ _2}{(1-\alpha_{\mathcal{A}})\epsilon}\sqrt{ \mathcal{I} } + n \mathcal{A}  +  \mathcal{A} ^2\right)$

Table 1: Summary of results. Throughout this work,  $n$  is the number of observations,  $d$  is the number of features,  $\mathcal{A}$  is the active set, and  $\mathcal{I}$  is the inactive set. In addition,  $\epsilon > 0$  is the approximation error,  $\delta \in (0, 1)$  is the failure probability,  $T$  is the number of kinks in the regularisation path  $\mathcal{P}$ ,  $\mathbf{X}_{\mathcal{A}} = [\mathbf{X}_i]_{i \in \mathcal{A}} \in \mathbb{R}^{n \times |\mathcal{A}|}$  is the matrix formed by the columns of  $\mathbf{X}$  in  $\mathcal{A}$ ,  $\mathbf{X}_{\mathcal{A}}^{\dagger}$  is the Moore–Penrose inverse of  $\mathbf{X}_{\mathcal{A}}$ , and the parameters  $\alpha_{\mathcal{A}}, \gamma_{\mathcal{A}} \in (0, 1)$  are defined in [Result 3](#). The notation  $\tilde{O}(\cdot)$  omits poly log terms in  $n, d, T$ , and  $\delta$ .

# Random Matrix





Algorithm	Error	Time complexity per iteration
Classical LARS 1	None	$O(nd)$
Simple quantum LARS 2	None	$\tilde{O}(n\sqrt{d} + n \mathcal{A} )$
Approximate classical LARS 4	$\lambda\epsilon\ \tilde{\beta}\ _1$	$\tilde{O}(d/\epsilon^2 + n \mathcal{A} )$
Approximate quantum LARS 3	$\lambda\epsilon\ \tilde{\beta}\ _1$	$\tilde{O}(\sqrt{d}/\epsilon + n \mathcal{A} )$

Table 2: Summary of results for the case when the design matrix  $\mathbf{X} \in \mathbb{R}^{n \times d}$  is a random Gaussian matrix with rows sampled i.i.d. from  $\mathcal{N}(\mathbf{0}, \Sigma)$  with well-behaved covariance matrix  $\Sigma > \mathbf{0} \in \mathbb{R}^{d \times d}$  and the vector of observations  $\mathbf{y} \in \mathbb{R}^n$  is such that  $\|\mathbf{y}\|_1/\|\mathbf{y}\|_2 = \text{poly log } n$ . These results are valid for active and inactive sets  $\mathcal{A}$  and  $\mathcal{I}$  of size  $|\mathcal{A}| = O(n/\log d)$  and  $|\mathcal{I}| = \Theta(d)$ . The notation  $\tilde{O}(\cdot)$  omits poly log terms in  $n$ ,  $d$ , and the failure probability  $\delta$ .

## NUS and NTU joint work - Thank you

1. Errors accumulate with quantizing subroutines.
2. Finding kinks optimally.
3. Difficulties in choosing classical algorithms and then quantizing.
4. Links to compressed sensing - open problems.
5. Paper discusses fast and slow rates of convergence.

# References

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